

APPLICATIONS OF ANALYTIC COMBINATORICS IN MATHEMATICAL BIOLOGY

(joint with H. Chang, M. Drmota, E. Y. Jin, and Y.-W. Lee)

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Department of Applied Mathematics
National Chiao Tung University



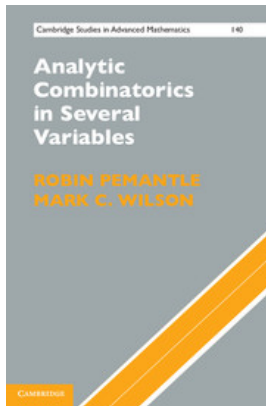
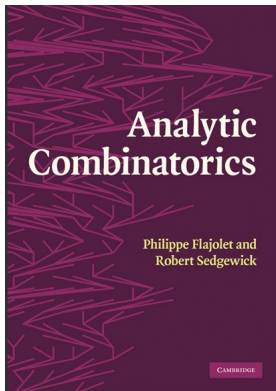
June 30th, 2015

Analytic Combinatorics (i)

Combinatorialists use recurrence, generating functions, and such transformations as the Vandermonde convolution; others, to my horror, use contour integrals, differential equations, and other resources of mathematical analysis.

- John Riordan (1968).

Analytic Combinatorics (ii)



Outline of the Talk

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- 2 Patterns in Phylogenetic Trees

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4 Number of Groups formed by Social Animals

M. Drmota, M. Fuchs, Y.-W. Lee (2016+). Stochastic analysis of the extra clustering model for animal grouping, in revision.

Evolutionary Biology



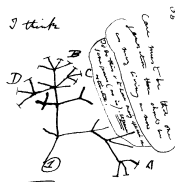
Charles Darwin
(1809-1882)

Evolutionary Biology



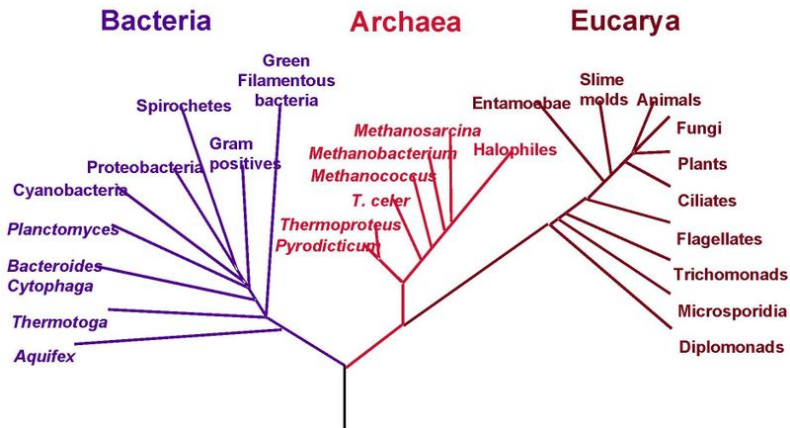
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First notebook on Transmutation of Species (1837)

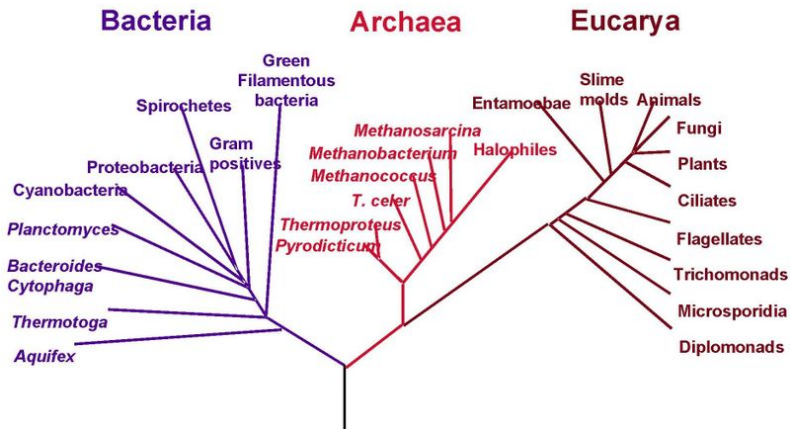


Then between A & B. various
size of union. C & B. The
first production, B & D
rather greater distance
Then forms would be
formed. - binary union

Phylogenetic trees (=PTs)



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Phylogenetic tree of size n : rooted, plane, unlabelled binary tree with n external nodes (and consequently $n - 1$ internal nodes).

Applications of PTs

- Understanding genetic relatedness of species

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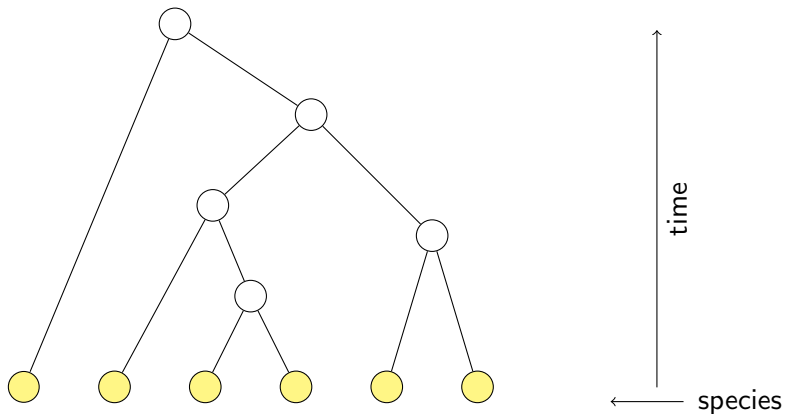
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Applications of PTs

- Understanding genetic relatedness of species
- Understanding the underlying evolutionary process
- Predicting possible future outcomes
- Testing appropriateness of random models
- Making conservation decisions in genetics
- Modeling the group formation process of social animals
- Etc.

Yule-Harding model or Kingman Coalescent

Example:



Yule-Harding model or Kingman Coalescent

Example:

Random Model 1:

At every time point,
two yellow nodes
uniformly coalescent.

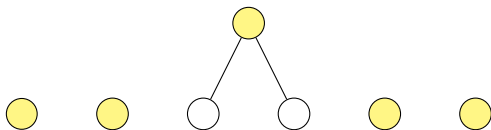


Yule-Harding model or Kingman Coalescent

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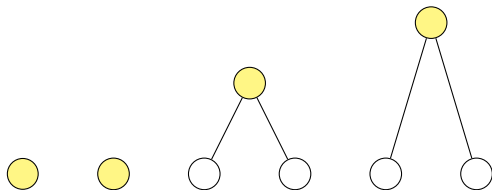
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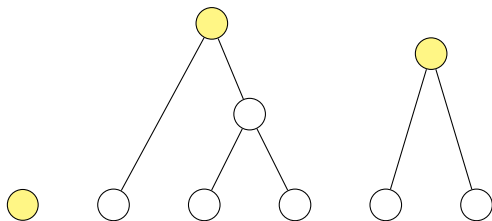


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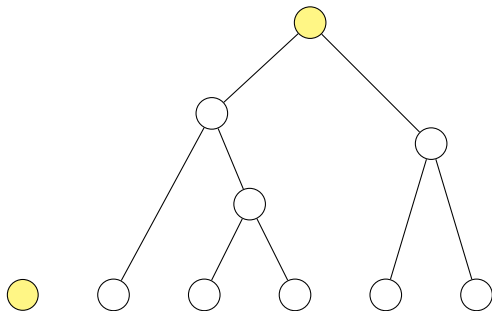


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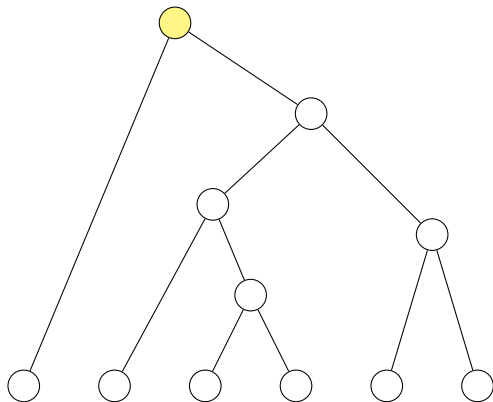


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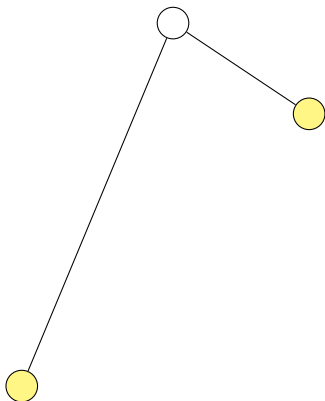


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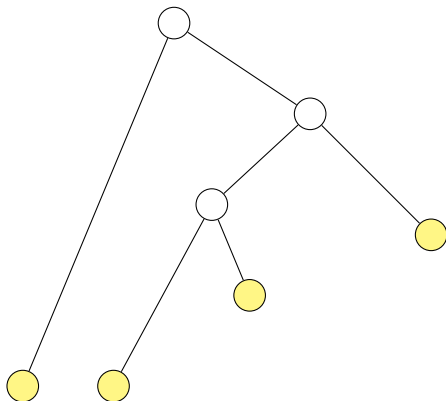


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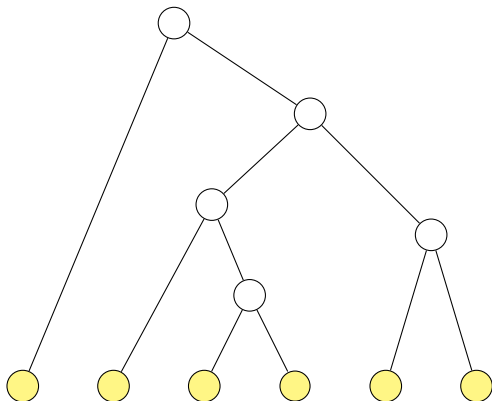


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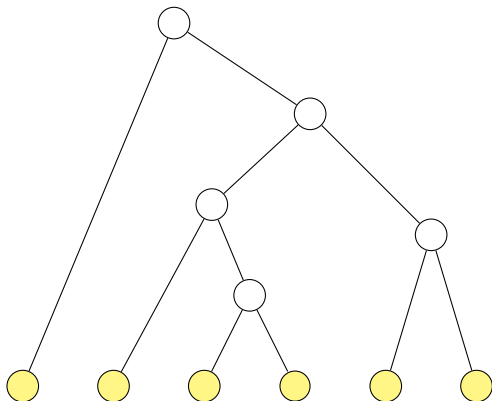


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Random model 1
and random model 2
are the same.

Binary Search Trees

Example: Input 1, 4, 2, 5, 3

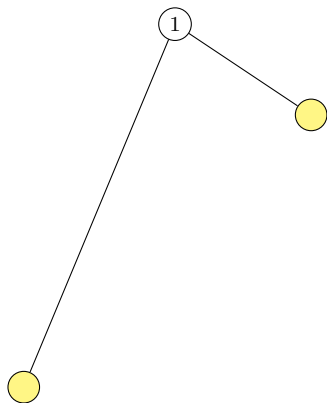


Random Model 3:

All permutations of the input are equally equally.

Binary Search Trees

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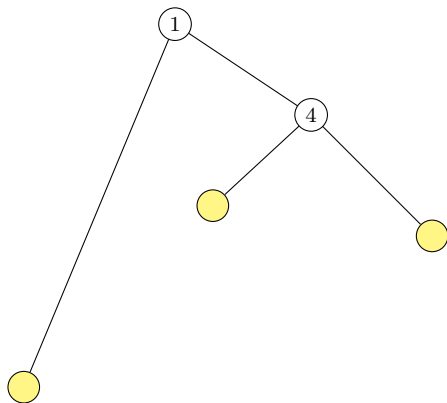


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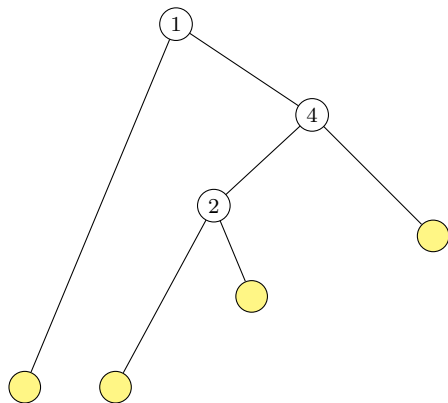


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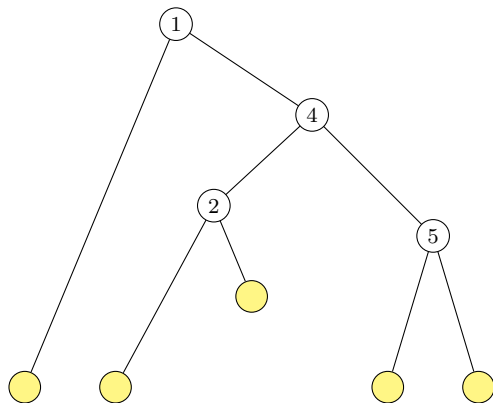


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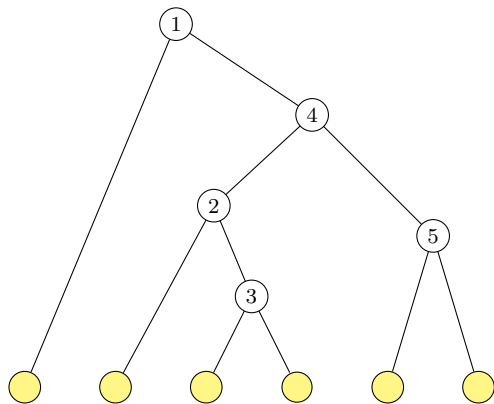


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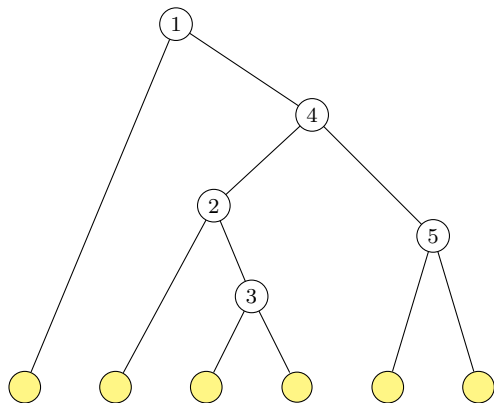


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Random Model 3:

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Again the same as random model 1 and random model 2.

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Node with an induced subtree of size k .

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Induced subtree of size k with an internal node which is descendent of all other internal nodes.

- Node with minimal clade size $k \geq 3$ (Blum and François (2005)):

Node with induced subtree of size k and either right or left subtree is an external node.

Mean and variance of k -pronged nodes

$X_{n,k}$ = # of k -pronged nodes in random PT of size n .

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Rosenberg 2006: we have

$$\mu_{n,k} := \mathbb{E}(X_{n,k}) = \frac{2n}{k(k+1)}, \quad (n > k)$$

and

$$\sigma_{n,k}^2 := \text{Var}(X_{n,k}) = \begin{cases} \frac{2(4k^2 - 3k - 4)(k-1)n}{k(k+1)^2(2k-1)(2k+1)}, & \text{if } n > 2k; \\ \frac{2(5k-7)(k-1)}{(k+1)^2(2k-1)}, & \text{if } n = 2k; \\ \frac{2(k^2 + k - 2n)n}{k^2(k+1)^2}, & \text{if } 2k > n > k. \end{cases}$$

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This result + central limit theorem also obtained by Devroye in **1991!**

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Question: What happens for $k \rightarrow \infty$ as $n \rightarrow \infty$?

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Theorem (Feng, Mahmoud, Panholzer; 2008)

(i) (Normal range) Let $k = o(\sqrt{n})$. Then,

$$\frac{X_{n,k} - \mu_{n,k}}{\sigma_{n,k}} \xrightarrow{d} \mathcal{N}(0, 1).$$

(ii) (Poisson range) Let $k \sim c\sqrt{n}$. Then,

$$X_{n,k} \xrightarrow{d} \text{Po}(2c^{-2}).$$

(iii) (Degenerate range) Let $k < n$ and $\sqrt{n} = o(k)$. Then,

$$X_{n,k} \xrightarrow{L_1} 0.$$

A General Framework

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We have,

$$X_{n,k} \stackrel{d}{=} X_{I_n,k} + X_{n-I_n,k}^*$$

where $X_{k,k} = \text{Bernoulli}(p_k)$, $X_{I_n,k}$ and $X_{n-I_n,k}^*$ are conditionally independent given I_n , and $I_n = \text{Unif}\{1, \dots, n-1\}$

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Here,

p_k	shape parameter
1	# of k -pronged nodes
$2/(k-1)$	# of nodes with minimal clade size k
$2^{k-2}/(k-1)!$	# of k caterpillars

Mean Value and Variance

We have

$$\mu_{n,k} = \frac{2p_k n}{k(k+1)}, \quad (n > k),$$

and

$$\sigma_{n,k}^2 = \frac{2(4k^3 + 4k^2 - k - 1 - (11k^2 - 5)p_k)p_k n}{k(k+1)^2(2k-1)(2k+1)}$$

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Note that

$$\mu_{n,k} \sim \sigma_{n,k}^2 \sim \frac{2p_k n}{k^2}$$

for $n > 2k$ and $k \rightarrow \infty$.

Poisson Approximation

Theorem (Chang and F.; 2010)

Let $k < n$ and $k \rightarrow \infty$. Then,

$$d_{TV}(X_{n,k}, \text{Po}(\mu_{n,k})) = \begin{cases} \mathcal{O}(p_k/k), & \text{if } \mu_{n,k} \geq 1; \\ \mathcal{O}(p_k/k \cdot \mu_{n,k}), & \text{if } \mu_{n,k} < 1. \end{cases}$$

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Recently re-proved by Holmgren and Janson with Stein's method:

C. Holmgren and S. Janson (2015). Limit laws for functions of fringe trees for binary search trees and recursive trees, *Electronic J. Probability*, **20:4**, 1–51.

Berry-Esseen bound and LLT

Theorem (Chang and F.; 2010)

For $\mu_{n,k} \rightarrow \infty$,

$$\sup_{x \in \mathbb{R}} \left| P \left(\frac{X_{n,k} - \mu_{n,k}}{\sigma_{n,k}} < x \right) - \Phi(x) \right| = \mathcal{O} \left(\frac{k}{\sqrt{p_k n}} \right).$$

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Theorem (Chang and F.; 2010)

For $\mu_{n,k} \rightarrow \infty$,

$$P(X_{n,k} = \lfloor \mu_{n,k} + x\sigma_{n,k} \rfloor) = \frac{e^{-x^2/2}}{\sqrt{2\pi}\sigma_{n,k}} \left(1 + \mathcal{O} \left((1 + |x|^3) \frac{k}{\sqrt{p_k n}} \right) \right),$$

uniformly in $x = o((p_k n)^{1/6}/k^{1/3})$.

Short Summary

We have proved:

- Explicit expressions for mean and variance.
- Local limit theorem + Berry-Esseen bound for the largest possible range of k .
- Poisson approximation + rate whenever $k \rightarrow \infty$.

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So, for $k \ll n$ the number of occurrences of the pattern can be approximated by the normal distribution, whereas for the remaining range a Poisson random variable should be used.

In particular, the phase change occurs much earlier than predicted by Feng, Mahmoud, and Panholzer.

Lloyd Shapley



Lloyd Shapley
(1923-)



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Shapley value:

Measure of importance of each player in a cooperative game



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→ recently used as prioritization tool of taxa in phylogenetics

Shapley Value and Modified Shapley Value

T ... PT;

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Shapley value $SV_T(a)$:

$$SV_T(a) = \frac{1}{n!} \sum_{S, a \in S} (|S| - 1)!(n - |S|)!(PD_T(S) - PD_T(S \setminus \{a\})).$$

Modified Shapley value $\widetilde{SV}_T(a)$:

$$\widetilde{SV}_T(a) = \frac{1}{n!} \sum_{|S| \geq 2, a \in S} (|S| - 1)!(n - |S|)!(PD_T(S) - PD_T(S \setminus \{a\})).$$

$PD(S)$ is the size of the ancestor of S .

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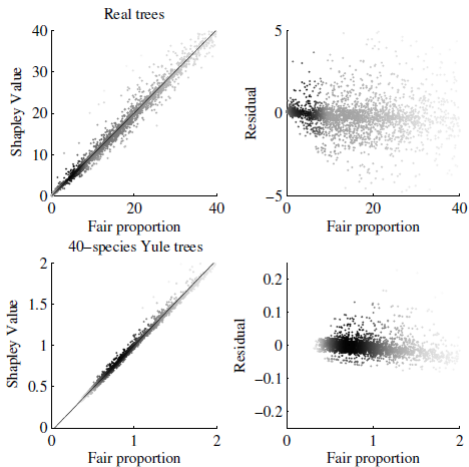
FP_n = fair proportion index of random taxon in random PT of size n :

$$FP_n | (I_n = j) = \begin{cases} \frac{1}{j} + FP_j, & \text{with probability } j/n; \\ \frac{1}{n-j} + FP_{n-j}, & \text{with probability } (n-j)/n, \end{cases}$$

where $I_n = \text{Unif}\{1, \dots, n-1\}$.

Strong Correlation between \widetilde{SV} and FP

Hartmann (2013):



SV = FP

Assume a is in left subtree T_l and $|T_l| = j$.

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Lemma

We have,

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and

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Theorem (F. and Jin; 2015)

We have,

$$SV_T(a) = FP_T(a).$$

\widetilde{SV} and FP (i)

$D_T(a)$ = depth of a in T :

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D_n = depth of random taxon in random PT of size n .

Lemma

We have,

$$\text{Var}(FP_n) = 10 - 6H_{n-1}^{(2)} - \frac{6}{n} - \frac{4}{n^2} \sim 10 - \pi^2$$

$$\text{Var}(D_n) = 2H_n - 4H_n^{(2)} + 2 \sim 2 \log n,$$

where $H_n = \sum_{1 \leq j \leq n} 1/j$ and $H_n^{(2)} = \sum_{1 \leq j \leq n} 1/j^2$.

\widetilde{SV} and FP (ii)

Lemma

We have,

$$\text{Cov}(\text{FP}_n, D_n) = 4H_{n-1}^{(2)} - 6 + \frac{2}{n} + \frac{4}{n^2} \sim \frac{2\pi^2}{6} - 6.$$

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Theorem (F. and Jin; 2015)

The correlation coefficient $\rho(\widetilde{SV}_n, \text{FP}_n)$ of modified Shapley value and fair proportion index tends to 1, i.e.,

$$\lim_{n \rightarrow \infty} \rho(\widetilde{SV}_n, \text{FP}_n) = 1.$$

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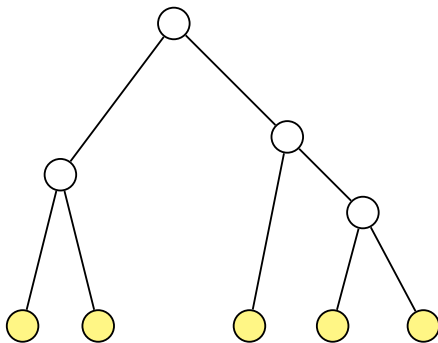
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- Fair proportion index was used by Zoological Society of London but its biodiversity value was unclear.
- Strong correlation between modified Shapley value and fair proportion index was observed by Hartmann and others.
- We proved that Shapley value equals fair proportion index.
- We showed that correlation coefficient of modified Shapley value and fair proportion index tends to 1 in Yule-Harding model and other random models.

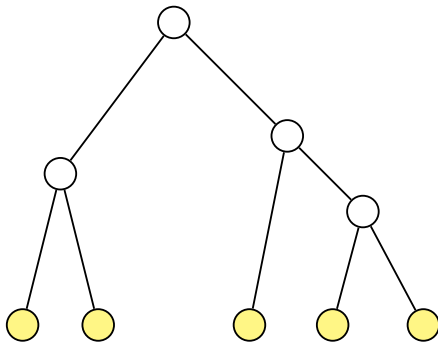
Social Animals

Consider n animals of a class of social animals. Construct a random PT with leaves representing the animals.



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Describes the genetic relatedness of the animals.

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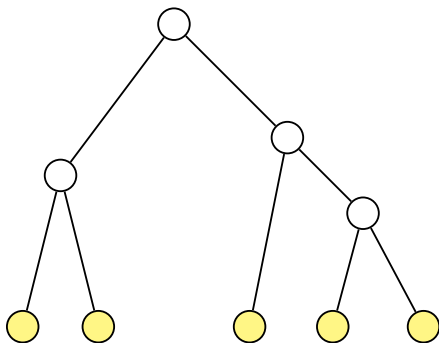
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Clade of a leaf:

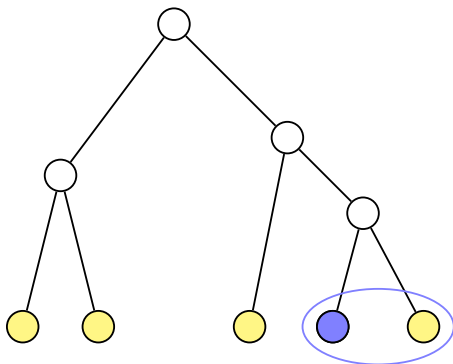
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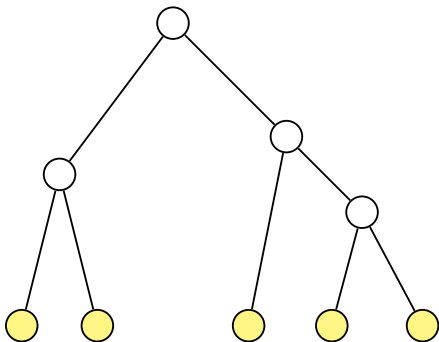
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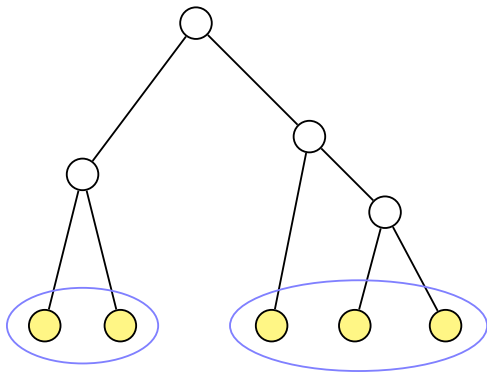
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2

of Groups

$X_n = \#$ of groups under the Yule Harding model

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We have,

$$X_n \stackrel{d}{=} \begin{cases} 1, & \text{if } I_n = 1 \text{ or } I_n = n - 1, \\ X_{I_n} + X_{n-I_n}^*, & \text{otherwise,} \end{cases}$$

where $I_n = \text{Uniform}\{1, \dots, n - 1\}$ is the # of animals in the left subtree and X_n^* is an independent copy of X_n .

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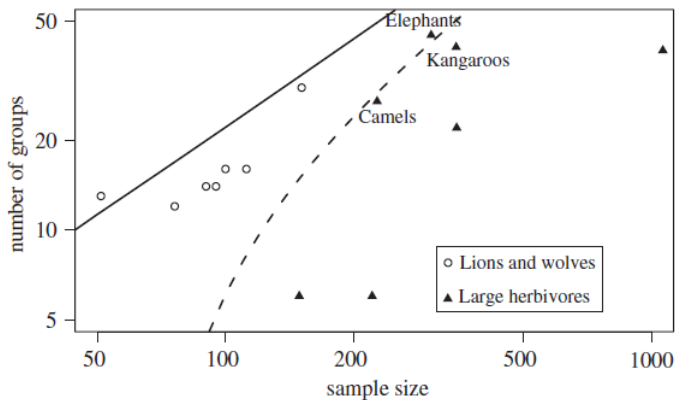
Theorem (Durand and François; 2010)

We have,

$$\mathbb{E}(X_n) \sim an \quad \left(a := \frac{1 - e^{-2}}{4} \right).$$

Comparison with Real-life Data

Durand, Blum and François (2007) presented the following data:



Group Patterns of Social Animals under the Neutral Model

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Department of Applied Mathematics,
National Chiao Tung University

This thesis was supervised by Dr. Michael Fuchs

May 26, 2012

Variance and SLLN

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We have,

$$\text{Var}(X_n) \sim \frac{(1 - e^{-2})^2}{4} n \log n = 4a^2 n \log n.$$

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We have,

$$P \left(\lim_{n \rightarrow \infty} \left| \frac{X_n}{\mathbb{E}(X_n)} - 1 \right| = 0 \right) = 1.$$

For SLLN, X_n is constructed on the same probability space via the tree evolution process underlying the Yule-Harding model.

Method of Moments

Theorem

Assume that $\mathbb{E}(X_n^k) \rightarrow \mathbb{E}(X^k)$ for all $k \geq 1$ and that X is uniquely characterised by its sequence of moments. Then,

$$X_n \xrightarrow{d} X.$$

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Many sufficient conditions for X being uniquely characterized by its moments are known, e.g.,

$$\sum_{k \geq 1} \mathbb{E}(X^k) \frac{z^k}{k!}$$

has a positive radius of convergence.

Higher Moments

Theorem (Lee; 2012)

For all $k \geq 3$,

$$\mathbb{E}(X_n - \mathbb{E}(X_n))^k \sim (-1)^k \frac{2k}{k-2} a^k n^{k-1}.$$

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Question: Is there a limit distribution?

Random Recursive Trees

Unordered, rooted trees.

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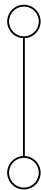
Uniformly choose one of the nodes and attach a child.



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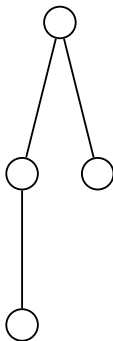
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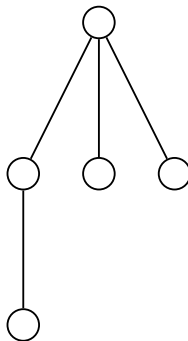
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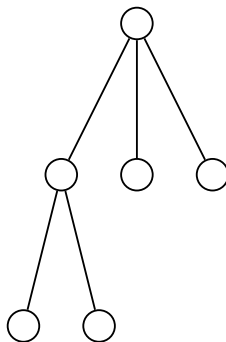
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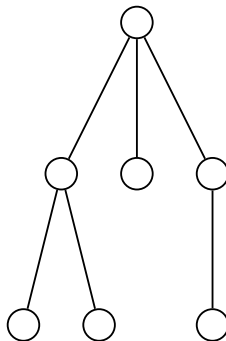
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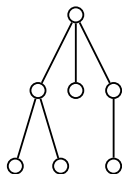
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Cutting Down Random Recursive Trees

Meir and Moon (1974):

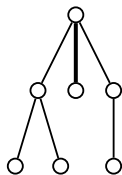
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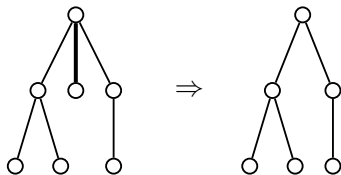
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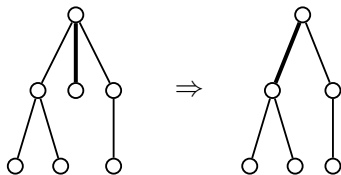
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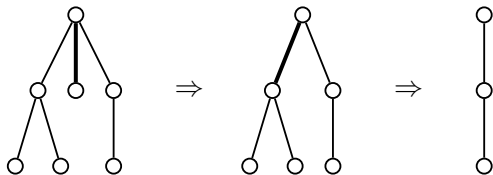
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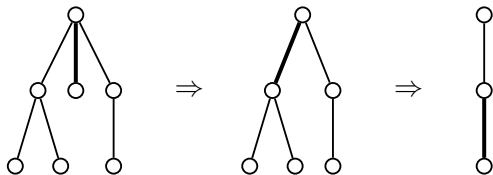
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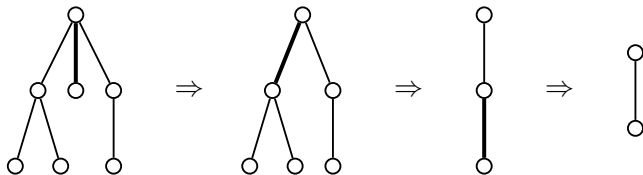
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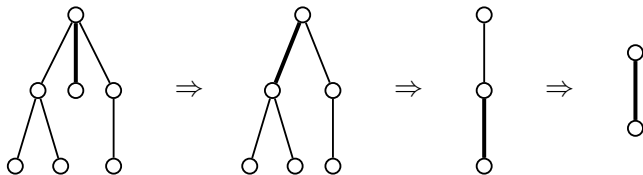
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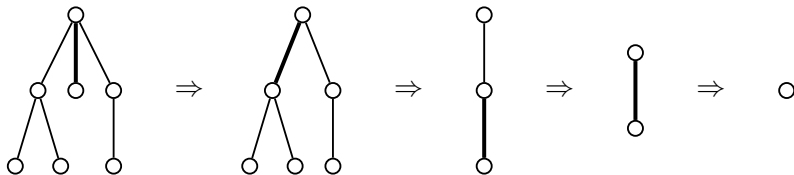
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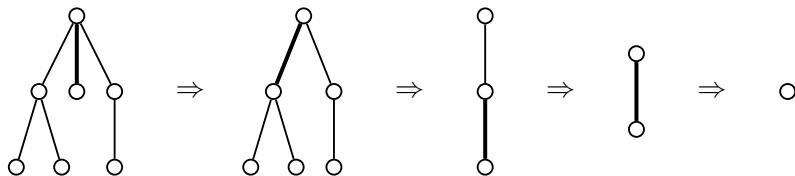
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$Y_n =$ number of steps until tree is destroyed
 $=$ number of edges cut $= 4$.

Mean, Variance and Higher Moments

Theorem (Panholzer; 2004)

We have,

$$\mathbb{E}(Y_n) \sim \frac{n}{\log n}$$

and for $k \geq 2$

$$\mathbb{E}(Y_n - \mathbb{E}(Y_n))^k \sim \frac{(-1)^k}{k(k-1)} \cdot \frac{n^k}{\log^{k+1} n}.$$

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Thus, again the limit law of

$$\frac{Y_n - \mathbb{E}(Y_n)}{\sqrt{\text{Var}(Y_n)}}$$

cannot be obtained from the method of moments!

Limit Law

Theorem (Drmotá, Iksanov, Moehle, Roessler; 2009)

We have,

$$\frac{\log^2 n}{n} Y_n - \log n - \log \log n \xrightarrow{d} Y$$

with

$$\mathbb{E}(e^{i\lambda Y}) = e^{i\lambda \log |\lambda| - \pi |\lambda|/2}.$$

The law of Y is spectrally negative stable with index of stability 1.

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Different proofs of this result exist.

Limit Law of X_n

Theorem (Drmotá, F., Lee; 2014)

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A probabilistic proof explaining the curious normalization was given recently.

S. Janson (2015). Maximal clades in random binary search trees, *Electronic J. Combinatorics*, **22**:1, paper 31.

Some Ideas of the Proof (i)

Set

$$X(y, z) = \sum_{n \geq 2} \mathbb{E}(e^{yX_n}) z^n.$$

Then,

$$z \frac{\partial}{\partial z} X(y, z) = X(y, z) + X^2(y, z) + e^y z^2 \frac{2e^y z^3}{1-z}.$$

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Some Ideas of the Proof (ii)

Set

$$\tilde{X}(y, z) = -\frac{V'(y, z)}{V(y, z)}.$$

Then,

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Solution is given by

$$V(y, z) = M_{-ey/2, 1/2} \left(2e^{y/2}(z-1) \right) + c(y)W_{-ey/2, 1/2} \left(2e^{y/2}(z-1) \right),$$

where

$$c(y) = -\frac{(e^{y/2} - 1) M_{-ey/2+1, 1/2} (-2e^{y/2})}{W_{-ey/2+1, 1/2} (-2e^{y/2})}.$$

Some Ideas of the Proof (iii)

Lemma

$V(y, z)$ is analytic in

$$\Delta = \{z \in \mathbb{C} : |z| < 1 + \delta\} \setminus \{\text{branch cut from } 1 \text{ to } \infty\}$$

for all $|y| < \eta$.

Moreover, $V(y, z)$ has only one (simple) zero with

$$z_0(y) = 1 - ay + 2a^2y^2 \log y + \mathcal{O}(y^2).$$

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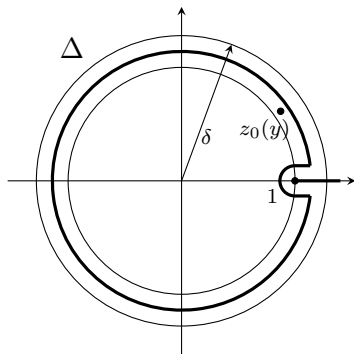
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$$\mathbb{E}(e^{yX_n}) = \exp\left(\frac{it\sqrt{n}}{2\sqrt{\log n}} - \frac{t^2}{4}\right) \left(1 + \mathcal{O}\left(\frac{\log \log n}{\log n}\right)\right).$$

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- More input from biologists is needed to make our results more relevant from the point of view of applications.