

THE NODE PROFILE OF SYMMETRIC DIGITAL SEARCH TREES

(joint with M. Drmota, H.-K. Hwang and R. Neinger)

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Department of Applied Mathematics
National Chiao Tung University



June 8th, 2015

Node Profile of (Rooted) Trees

$B_{n,k}$ = number of external nodes at level k ;

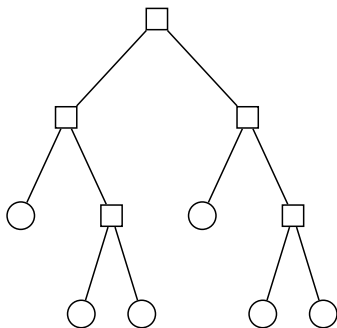
$I_{n,k}$ = number of internal nodes at level k .

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Example:

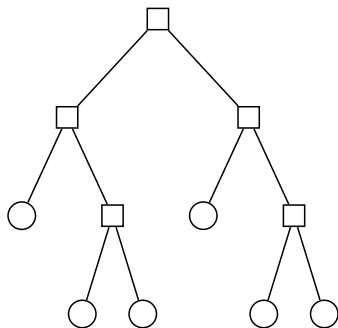


Node Profile of (Rooted) Trees

$B_{n,k}$ = number of external nodes at level k ;

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Example:



$$B_{5,0} = 0,$$

$$B_{5,1} = 0,$$

$$B_{5,2} = 2,$$

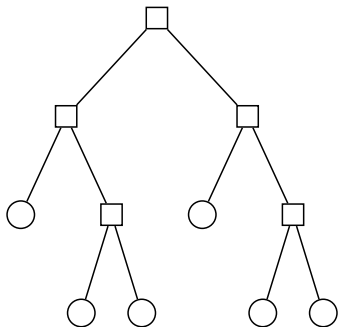
$$B_{5,3} = 4,$$

Node Profile of (Rooted) Trees

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$$B_{5,0} = 0, \quad I_{5,0} = 1;$$

$$B_{5,1} = 0, \quad I_{5,1} = 2;$$

$$B_{5,2} = 2, \quad I_{5,2} = 2;$$

$$B_{5,3} = 4, \quad I_{5,3} = 0.$$

Relations to Other Shape Parameters

Many shape parameters can be analyzed through the profile.

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- Depth: $P(D_n = k) = B_{n,k}/(n + 1)$;
- Width: $\max\{B_{n,k} : k \geq 0\}$;
- Total Path Length: $\sum_k k B_{n,k}$;
- Height: $\max\{k : B_{n,k} > 0\}$;
- Shortest Path: $\min\{k : B_{n,k} > 0\}$;
- Fill-up Level: $\max\{k : I_{n,k} = 2^k\}$;
- Etc.

Profile of Random Trees

- \sqrt{n} -Trees:

Aldous (1991); Drmota and Gittenberger (1997); Kersting (1998); Pitman (1999); etc.

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- $\log n$ -Trees:

- Binary Search Trees: Chauvin, Drmota, Jabbour-Hattab (2001); Drmota and Hwang (2005); F., Hwang, Neininger (2006).
- Recursive Trees: Drmota and Hwang (2005); F., Hwang, Neininger (2006).
- Plane-oriented Recursive Trees: Hwang (2007).
- m -ary Search Trees: Drmota, Janson, Neininger (2008).

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René de la Briandais (1959)

Name from data **re**trieval (suggested by Fredkin).

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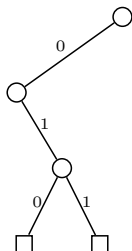
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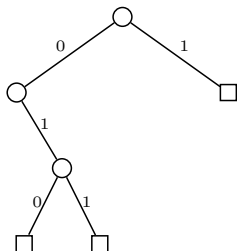
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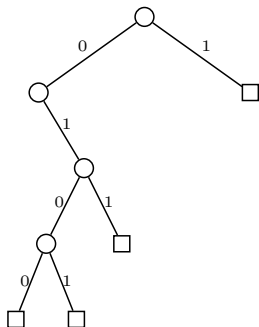
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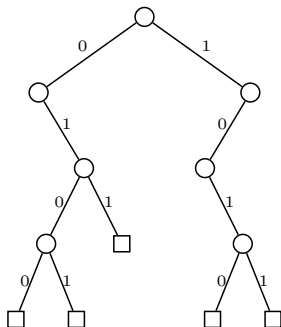
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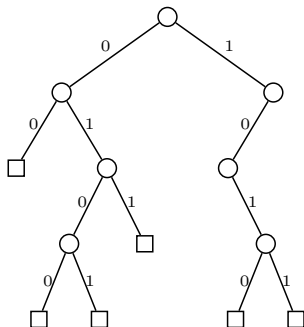
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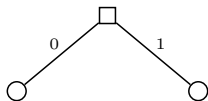
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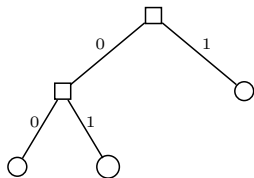
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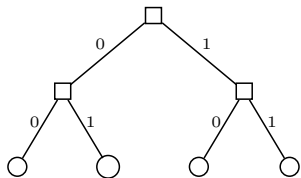
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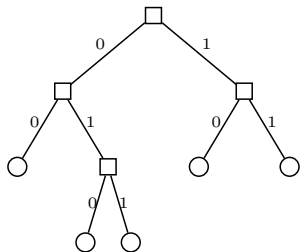
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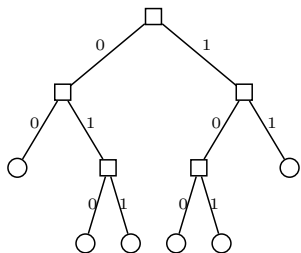
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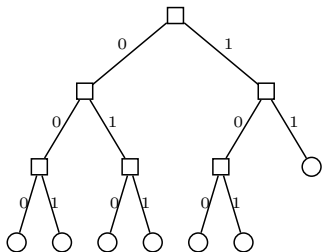
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Question: What can be said about the profile?

In this talk, we are interested in mean, variance and limit laws of the profile for symmetric DSTs.

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- Symmetric DSTs:

Variance & limit laws: Drmota, F., Hwang, Neininger (→ this talk).

Profile of Tries

Hwang, Nicodème, Park, Szpankowski (2009)

Profile of Tries

G. Park¹, H.-K Hwang², P. Nicodème³, and W. Szpankowski⁴

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² Institute of Statistical Science, Academia Sinica, 11529 Taipei, Taiwan
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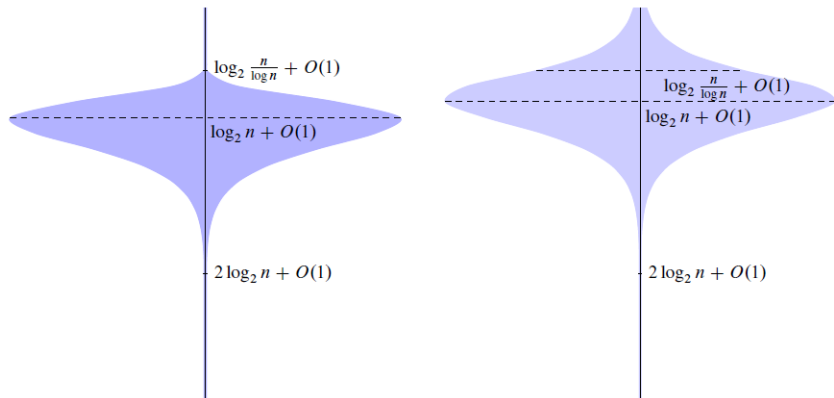
³ Laboratory LIX, École polytechnique, 91128 Palaiseau Cedex, France
nicodeme@lix.polytechnique.fr

⁴ Department of Computer Sciences, Purdue University, 250 N. University Street,
West Lafayette, Indiana, 47907-2066, USA
spa@cs.purdue.edu

Abstract. The *profile* of a trie, the most popular data structures on words, is a parameter that represents the number of nodes (either internal or external) with the same distance to the root. Several, if not

Plot of Mean Profile of Symmetric Tries

Hwang, Nicodéme, Park, Szpankowski (2009):



Symmetric Tries: Mean

We have,

$$\mu_{n,k} := \mathbb{E}(B_{n,k}) \sim \begin{cases} n(1 - 2^{-k})^{n-1}, & \text{if } 2^{-k}n \rightarrow \infty; \\ \tilde{M}_{k,1}(n), & \text{if } 4^{-k}n \rightarrow 0, \end{cases}$$

where

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$$\tilde{M}_{k,1}(n) \sim \begin{cases} ne^{-n/2^k}, & \text{if } 2^{-k}n \rightarrow \infty; \\ \Theta(n), & \text{if } 2^{-k}n = \Theta(1); \\ 2^{-k}n^2, & \text{if } 2^{-k}n \rightarrow 0. \end{cases}$$

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Thus, the profile has maximum of order n (asymmetric tries: $n/\sqrt{\log n}$)

Symmetric Tries: Variance

We have,

$$\sigma_{n,k}^2 := \text{Var}(B_{n,k}) \sim \begin{cases} n(1 - 2^{-k})^{n-1}, & \text{if } 2^{-k}n \rightarrow \infty; \\ \tilde{V}_k(n), & \text{if } 4^{-k}n \rightarrow 0, \end{cases}$$

where

$$\begin{aligned} \tilde{V}_k(z) = & z(e^{-z/2^k} - e^{-z/2^{k-1}}) + 2^{-k}z^2e^{-z/2^{k-1}} \\ & - 2^{1-k}z^2(e^{-z/2^k} - e^{-z/2^{k-1}})^2. \end{aligned}$$

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In particular,

$$\tilde{V}_k(n) \sim \begin{cases} ne^{-n/2^k} \sim \tilde{M}_k(n), & \text{if } 2^{-k}n \rightarrow \infty; \\ \Theta(n), & \text{if } 2^{-k}n = \Theta(1); \\ 2^{1-k}n^2 \sim 2\tilde{M}_k(n), & \text{if } 2^{-k}n \rightarrow 0. \end{cases}$$

Poissonization and Depoissonization

Poisson Model: Build digital tree from Poisson-distributed number of records.

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Poisson moments:

$$\tilde{M}_{k,\ell}(z) = \mathbb{E}(B_{\text{Pois}(z),k}^\ell) = e^{-z} \sum_{n \geq 0} \mathbb{E}(B_{n,k}^\ell) \frac{z^n}{n!}.$$

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Poisson Heuristic:

$$\tilde{M}_{k,\ell}(z) \text{ sufficiently smooth} \implies \mathbb{E}(B_{n,k}^\ell) \approx \tilde{M}_{k,\ell}(n).$$

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Poisson heuristic made precise by the **Theory of Analytic Depoissonization** (Jacquet & Szpankowski; 1998).

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With this choice:

$$\text{Var}(B_{n,k}) \sim \tilde{V}_k(n)$$

when $4^{-k}n \rightarrow 0$.

Symmetric DSTs: Mean

Let

$$Q(z) = \prod_{\ell=1}^{\infty} (1 - z2^{-\ell}), \quad Q_n = \prod_{\ell=1}^n (1 - 2^{-\ell}) = \frac{Q(2^{-n})}{Q(1)}.$$

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Theorem

We have,

$$\mu_{n,k} \begin{cases} \sim \frac{2^k}{Q_k} (1 - 2^{-k})^n, & \text{if } 2^{-k}n \rightarrow \infty; \\ = 2^k F(n/2^k) + \mathcal{O}(1), & \text{if } 4^{-k}n \rightarrow 0, \end{cases}$$

where $F(x)$ is the positive function

$$F(x) = \sum_{j \geq 0} \frac{(-1)^j 2^{-\binom{j}{2}}}{Q_j Q(1)} e^{-2^j x}.$$

$F(x)$ (i)

As $x \rightarrow \infty$,

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and as $x \rightarrow 0$,

$$F(x) \sim \frac{X^{1/\log 2}}{\sqrt{2\pi x}} \exp\left(-\frac{(\log X \log X)^2}{\log 2} - \sum_{j \in \mathbb{Z}} c_j (X \log X)^{-\chi_j}\right),$$

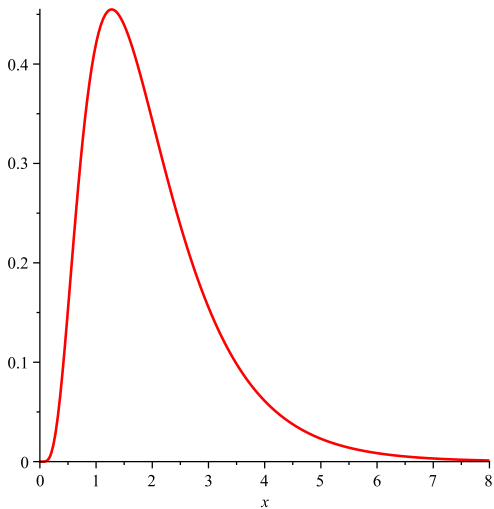
where $X = 1/(x \log 2)$, $\chi_j = 2j\pi i / \log 2$,

$$c_0 = \frac{\log 2}{12} + \frac{\pi^2}{6 \log 2}$$

and

$$c_j = \frac{1}{2j \sinh(2j\pi / \log 2)}, \quad (j \neq 0).$$

$F(x)$ (ii)



Some Details of the Proof (i)

We have,

$$\tilde{M}_{k,1}(z) + \tilde{M}'_{k,1}(z) = 2\tilde{M}_{k-1,1}(z/2).$$

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$$\tilde{M}_{k,1}(z) = 2^k \sum_{0 \leq j \leq k} \frac{(-1)^j 2^{-\binom{j}{2}}}{Q_j Q_{k-j}} e^{-z/2^{k-j}}.$$

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From this,

$$\mu_{n,k} = 2^k \sum_{0 \leq j \leq k} \frac{(-1)^j 2^{-\binom{j}{2}}}{Q_j Q_{k-j}} \left(1 - 2^{j-k}\right)^n.$$

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This is useful if $n2^{-k} \rightarrow \infty$.

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We have,

$$\tilde{M}_{k,1}(z) = 2^k \sum_{r \geq 0} \frac{2^{-\binom{r+1}{2} - kr}}{Q_r} F^{(r)} \left(\frac{z}{2^k} \right).$$

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This gives,

$$\tilde{M}_{k,1}(z) = 2^k F\left(\frac{z}{2^k}\right) + \mathcal{O}(1).$$

Result follows from depoissonization.

Symmetric DSTs: Variance

Theorem (Drmotá, F., Hwang, Neininger)

We have,

$$\sigma_{n,k}^2 \begin{cases} \sim \frac{2^k}{Q_k} \left(1 - 2^{-k}\right)^n, & \text{if } 2^{-k}n \rightarrow \infty; \\ = 2^k H(n/2^k) + \mathcal{O}(1), & \text{if } 4^{-k}n \rightarrow 0, \end{cases}$$

where $H(x)$ is a function with

$$H(x) = \frac{e^{-x}}{Q(1)} + \mathcal{O}(xe^{-2x}), \quad (x \rightarrow \infty)$$

and

$$H(x) \sim 2F(x), \quad (x \rightarrow 0).$$

$H(x)$ (i)

We have,

$$H(x) = \sum_{j,r=0}^{\infty} \sum_{0 \leq h, \ell \leq j} \frac{2^{-j} (-1)^{r+h+\ell} 2^{-\binom{r}{2} - \binom{h}{2} - \binom{\ell}{2} + 2h + 2\ell}}{Q_r Q(1) Q_h Q_{j-h} Q_\ell Q_{j-\ell}} \varphi(2^{r+j}, 2^h + 2^\ell; x),$$

where

$$\varphi(u, v; x) = \begin{cases} \frac{e^{-ux} - ((v-u)x + 1)e^{-vx}}{(v-u)^2}, & \text{if } u \neq v; \\ x^2 e^{-ux} / 2, & \text{if } u = v. \end{cases}$$

$H(x)$ (i)

We have,

$$H(x) = \sum_{j,r=0}^{\infty} \sum_{0 \leq h, \ell \leq j} \frac{2^{-j} (-1)^{r+h+\ell} 2^{-\binom{r}{2} - \binom{h}{2} - \binom{\ell}{2} + 2h + 2\ell}}{Q_r Q(1) Q_h Q_{j-h} Q_\ell Q_{j-\ell}} \varphi(2^{r+j}, 2^h + 2^\ell; x),$$

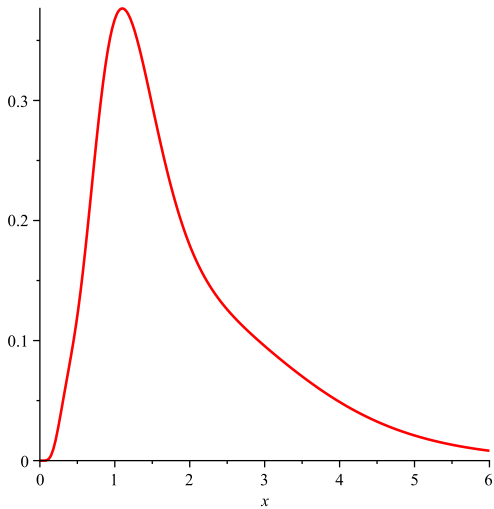
where

$$\varphi(u, v; x) = \begin{cases} \frac{e^{-ux} - ((v-u)x + 1)e^{-vx}}{(v-u)^2}, & \text{if } u \neq v; \\ x^2 e^{-ux} / 2, & \text{if } u = v. \end{cases}$$

Proposition (Drmota, F., Hwang, Neininger)

$H(x)$ is a positive function on $(0, \infty)$.

$H(x)$ (ii)



Some Details of the Proof (i)

We have,

$$\tilde{V}_k(z) + \tilde{V}'_k(z) = 2\tilde{V}_{k-1}(z/2) + z\tilde{M}''_{k,2}(z)^2.$$

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By Laplace transform and its inverse,

$$\tilde{V}_k(z) = \sum_{(j,r,h,\ell) \in \mathcal{V}} \frac{2^{k-j}(-1)^{r+h+\ell} 2^{-\binom{r}{2} - \binom{h}{2} - \binom{\ell}{2} + 2h + 2\ell}}{Q_r Q_{k-j-r} Q_h Q_{j-h} Q_\ell Q_{j-\ell}} \varphi\left(2^{r+j}, 2^h + 2^\ell, \frac{z}{2^k}\right)$$

with

$$\mathcal{V} = \{(j, r, h, \ell) : 0 \leq j \leq k, 0 \leq r \leq k - j, 0 \leq h, \ell \leq j\}$$

and

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Some Details of the Proof (ii)

Lemma

We have,

$$\tilde{V}_k(z) = 2^k \sum_{m \geq 0} \frac{2^{-\binom{m+1}{2} - km}}{Q_m} H^{(m)} \left(\frac{z}{2^k} \right).$$

Some Details of the Proof (ii)

Lemma

We have,

$$\tilde{V}_k(z) = 2^k \sum_{m \geq 0} \frac{2^{-(\binom{m+1}{2}) - km}}{Q_m} H^{(m)} \left(\frac{z}{2^k} \right).$$

The Laplace transform of $H(z)$:

$$\mathcal{L}[H(z); s] = \sum_{j \geq 0} 4^{-j} \frac{\tilde{g}_j^*(2^{-j}s)}{Q(-2^{1-j}s)}$$

where

$$\tilde{g}_j^*(s) = \sum_{0 \leq k, \ell \leq j} \frac{(-1)^{h+\ell} 2^{-\binom{h}{2} - \binom{\ell}{2} + 2h + 2\ell}}{Q_k Q_{j-k} Q_\ell Q_{j-\ell}} \frac{1}{(2^j s + 2^h + 2^\ell)^2}.$$

Some Details of the Proof (iii)

Lemma

We have, as $s \rightarrow \infty$,

$$\frac{\tilde{g}_0^*(s)}{Q(-2s)} \sim \frac{1}{s^2 Q(-2s)}, \quad 4^{-1} \frac{\tilde{g}_1^*(2^{-1}s)}{Q(-s)} \sim \frac{9}{sQ(-2s)}$$

and, for $j \geq 2$,

$$4^{-j} \frac{\tilde{g}_j^*(2^{-j}s)}{Q(-2^{1-j}s)} \sim \frac{(2j-3)!}{((j-2)!)^2} \frac{2^{\binom{j}{2}}}{s^{j-2} Q(-2s)}.$$

Thus,

$$\mathcal{L}[H(z); s] \sim \frac{2}{Q(-2s)}$$

and hence, $H(x) \sim 2F(x)$ as $x \rightarrow 0$.

Symmetric DSTs: Limit Laws

Corollary (Drmota, F., Hwang, Neininger)

We have,

$$\mu_{n,k} \longrightarrow \infty \quad \text{iff} \quad \sigma_{n,k}^2 \longrightarrow \infty.$$

Theorem (Drmota, F., Hwang, Neininger)

Assume that $\mu_{n,k} \longrightarrow \infty$. Then,

$$\frac{B_{n,k} - \mu_{n,k}}{\sigma_{n,k}} \xrightarrow{d} N(0, 1),$$

where $N(0, 1)$ denotes a standard normal distribution.

Application to the Height

H_n = height of a symmetric DST of size n .

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Theorem (Drmotá, F., Hwang, Neininger)

Set

$$k_n = \min\{k \geq \log_2 n : 2^k F(n/2^k) \leq 1\}.$$

Then,

$$k_n = \log_2 n + \sqrt{2 \log_2 n} - \log_2 \left(\sqrt{\log_2 n} \right) + \mathcal{O}(1).$$

Moreover,

$$P(H_n = k_n - 2 \text{ or } H_n = k_n - 1) \rightarrow 1.$$

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This solves an open problem of Aldous & Shields.

Summary of Results for Symmetric DSTs

- Mean profile tends to infinity when k is roughly in the range

$$\log_2 n - \log_2 \log n \leq k \leq \log_2 n + \sqrt{2 \log_2 n};$$

otherwise it is bounded.

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- If mean tends to infinity, a central limit theorem holds.
- Our results have many applications, e.g., they allow us to solve a problem of Aldous & Shields.